



Article Cut Points of the Conicity Index and Associated Factors in Brazilian Rural Workers

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Abstract: (1) Background: Metabolic syndrome is associated with cardiovascular complications. Therefore, this study aims to establish cut points for the conicity index based on the components of metabolic syndrome and to associate it with characteristic sociodemographic, food consumption and occupational factors in Brazilian rural workers; (2) Methods: A cross-sectional study carried out with farmers. The receiver operating characteristic curve was calculated and the cut-off points for the conicity index were identified by the area under the curve, sensitivity and specificity. The variables included in the binary logistic regression analysis were selected by considering p < 0.20 in the bivariate test; (3) Results: The cut points were similar in females according to both criteria, resulting in a single cut-off of 1.269. In males, the cut points showed differences, resulting in 1.272 according to the NCEP-ATP III and 1.252 according to the IDF. We have shown that younger people, those who work more than 40 h a week and the lowest contribution of culinary ingredients are associated with increased odds of abdominal obesity, while the consumption of the products they sell or produce decreases these chances; (4) Conclusions: The conicity index showed high discriminatory power for the identification of abdominal obesity in rural workers. Therefore, there is a need to improve eating habits and promote healthier eating environments for individuals, respecting traditional food culture, mainly to contain the advance of MS in rural areas.

Keywords: rural health; abdominal obesity; metabolic syndrome

1. Introduction

Metabolic syndrome (MetS) is defined by a set of abnormalities, such as arterial hypertension, dyslipidemia, insulin resistance and obesity [1]. In Brazil, a multicenter epidemiological survey recorded a prevalence of MetS of 38.4%, with a high proportion in women, elderly individuals and less educated individuals [2]. However, studies on the health of rural workers are still scarce in the scientific literature, with those related to conditions intrinsic to work in the field being more predominant. This, added to the difficult access to primary health care and lower availability of services and specialized professionals, ends up generating underreporting of diagnosed diseases and mortality rates [3].



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Copyright: © 2022 by the authors. Licensee MDPI, Basel, Switzerland. This article is an open access article distributed under the terms and conditions of the Creative Commons Attribution (CC BY) license (https:// creativecommons.org/licenses/by/ 4.0/). A recent study by Luz et al. [4] showed that six out of ten farmers in a region of southeastern Brazil had at least one cardiovascular risk factor, with arterial hypertension and dyslipidemia being the most prevalent. It is also worth noting that being overweight more than doubled the farmers' chances of having two or more cardiovascular risk factors. In this same population, a 33.8% prevalence of insulin resistance was found, with obesity or being overweight increasing the risk of an individual presenting this condition by about three times [5].

The role of obesity as a potential metabolic risk factor is already well defined in the literature [6–10]. In a recent study with Brazilian farmers, weight was assessed from the body mass index (BMI) and was found to be higher than the general population [11]. It is known that BMI is an anthropometric indicator widely used to categorize nutritional status, however, it is limited by the fact that it does not provide the distribution of body fat [12–14]. Thus, the conicity index appears as an alternative measure to assess abdominal obesity [15].

The performance of the conicity index as an indicator of abdominal obesity from the MetS components has been demonstrated in numerous studies, but there is still no consensus on the cut points, which vary according to population, sex and age [1–17]. Until then, no cut point has been determined for this index in the Brazilian rural population, which justifies the importance of the present study.

Considering the risks to which rural workers are exposed, the difficulty in accessing health services and the high prevalence of multimorbidity in this population [18], it is necessary to establish a cut point for the index of conicity in order to generate an instrument capable of discriminating abdominal fat that can be used not only as a support material for policies and programs aimed at the health of rural workers, but also as a tool to be used by health professionals in the routines involved in primary care in rural areas.

Therefore, this study aims to establish cut points for the conicity index based on the components of metabolic syndrome according to the International Diabetes Federation (IDF) and the National Cholesterol Education Program's Adult Treatment Panel III (NCEP-ATP III), and to associate it with characteristic sociodemographic, food consumption and occupational factors in Brazilian rural workers.

2. Materials and Methods

2.1. Study Design and Population

This is a cross-sectional study included in a larger project entitled "Health condition and associated factors: a study with farmers in Espírito Santo—AgroSaúdES", funded by the Fundação de Amparo à Pesquisa do Espírito Santo (FAPES). The study was approved by the Research Ethics Committee of the Health Sciences Center of the Federal University of Espírito Santo, opinion n° 2091172 (CAAE 52839116.3.0000.5060) and meets the requirements required by Resolution of the National Health Council n. 466/12 and its complements for research involving human beings.

The study was carried out in the municipality of Santa Maria de Jetibá located in the state of Espírito Santo, southeastern Brazil, and it had a representative sample of farmers who met the following inclusion criteria: age between 18 and 59 years, not being pregnant, having agriculture as the main source of income and being in full employment for at least six months.

In the sample calculation, the population of farmers in the region (N = 7287), the expected prevalence of abdominal obesity of 50% and a significance level of 5% were considered. The minimum sample size was 365. However, as a way of improving the representativeness of the sample, data from 781 farmers who were investigated in the original project and who answered questions about the investigated outcome were analyzed.

Data collection took place between December 2016 and April 2017 on the premises of the municipality's health units by properly trained researchers. The details involved in the data collection and research development are detailed in the article by Petarli et al. [18].

2.2. Data Collection

Farmers underwent blood collection for biochemical tests. As a way of minimizing errors, the analyzes were performed by a single laboratory and the fasting time required to perform the exams was 12 h.

The determination of HDL cholesterol was determined by the enzymatic colorimetric method with the Cholesterol Liquicolor Kit (In Vitro Diagnostica Ltd.a, Belo Horizonte, MG, Brazil) and the Cholesterol HDL Precipitation Kit (In Vitro Diagnostica Ltd.a, Belo Horizonte, MG, Brazil). Triglycerides were determined by the enzymatic colorimetric method with the Liquicolor mono[®] Triglycerides Kit (In Vitro Diagnostica Ltd.a, Belo Horizonte, MG, Brazil).

Systolic blood pressure (SBP) and diastolic blood pressure (DBP) were measured during the interview 3 times for each subject using the Omron[®] HME-7200 Automatic Pressure Monitor, calibrated and validated by the National Institute of Metrology, Quality and Technology (INMETRO). In order not to interfere with the results, the individuals were instructed to remain seated and rest for about 5 min, empty their bladders and not consume food, alcohol, coffee or cigarettes in the 30 min prior to the assessment. For data analysis, the average of two measurements was considered and a third measurement was performed whenever the difference between the first two was greater than 4 mmHg [18,19].

Food consumption data were obtained by applying three 24-h recalls (R24h) during the interview, two days from the week and one day from the weekend within 15 days after the first R24h in the return interviews. After the registration of the food and acquisition of the calories, no exclusion was performed due to extremes in energy consumption [20]. After obtaining the values of each R24h, the analysis of the attenuation was performed using the PC-SIDE software (Department of Statistics, Iowa State University, Ames, IA, USA), which follows the methodology of Nusser et al. [21]. Then, the adjustment for energy by the residual method was carried out, which corrects the estimates of nutrients by total energy intake [20]. The foods consumed by farmers in the three R24h were listed (n = 355 food items). These foods were classified according to NOVA, that is, in four groups: in natura and minimally processed foods; processed culinary ingredients; processed foods; and ultra-processed foods [22–24]. After the classification of food items in each NOVA group, the calories from each food group and subgroup were added. Then, the caloric contribution of each food group to daily energy consumption was calculated [23,25]. The energy contribution of each food group was categorized as "lower contribution" when below the median and "higher contribution" when above the median. More information about this collection and classification can be found in Cattafesta et al. [26].

Weight measurement was performed with participants barefoot, in an upright position, wearing as little clothing as possible [27] using a portable Tanita[®] scale. Height was measured with the individuals barefoot, standing, in an upright position, arms extended along the body and eyes fixed on a point on the horizon [27] using a portable Sanny[®] stadiometer. Waist circumference (WC) was measured with the participant standing, with arms extended along the body and feet together, with the inelastic tape positioned at the midpoint between the lower edge of the costal arch and the iliac crest [27]. Three non-consecutive measurements were taken, the first being discarded and the average of the last two considered as the final measurement. In relation to BMI, it was categorized considering the WHO cut points [28] in eutrophy/low weight and overweight/obesity.

The conicity index was calculated from the measurements of body mass, height and WC through the following mathematical equation [15]:

Conicity Index =
$$\frac{\text{Waist circumference (m)}}{0.109\sqrt{\frac{\text{Body mass (kg)}}{\text{Stature (m)}}}}$$

To establish the index cut-off point for both sexes, the MS criteria according to the IDF and NCEP-ATP III were used. In the first, the individual with abdominal obesity

is considered to have the syndrome, assessed by a WC \geq 84 cm for women or \geq 94 cm for men, occurring simultaneously with two more criteria: fasting blood glucose \geq 100; SBP \geq 130 mmHg or DBP \geq 85 mmHg; TG \geq 150 mg/dL and HDL-c < 40 mg/dL for men and <50 mg/dL for women. The use of antihypertensive, hypoglycemic and/or lipid-lowering drugs are considered in both criteria for MS, as they classify the individual with hypertension, diabetes and/or dyslipidemia, respectively [29]. The NCEP considers MS in the presence of at least three of the following criteria: WC > 102 cm for men or >88 cm for women; HDL-c < 40 mg/dL for men and <50 mg/dL for women; TG \geq 150 mg/dL; SBP \geq 130 mmHg and DBP \geq 85 mmHg and fasting blood glucose \geq 100.

2.3. Statistical Analysis

To describe the study variables, absolute and relative frequencies were used. To verify if there was a difference between the proportions of the independent variables and the outcome, Pearson's chi-square test (x^2) was used for qualitative variables.

The data were submitted to analysis of the ROC curve (Receiver Operating Characteristic) to establish the cut points for the conicity index according to the set of conditions that make up the criteria for MetS in both diagnostic criteria mentioned above. Cut points were defined based on accuracy, specificity and sensitivity.

A hierarchical logistic regression was performed for the association of the conicity index with the independent variables, including the variables that presented a value of $p \le 0.20$ in the bivariate analysis. From the first to the fourth models, socioeconomic, behavioral, food consumption, anthropometric and work variables were aggregated, respectively. Among the sociodemographic variables evaluated were sex, age group ("up to 29 years", "30 to 39 years", "40 to 49 years" and "50 years or more"), schooling ("less than 4 years", "4 to 8 years" and "more than 8 years"), marital status ("unmarried", "married or lives with a partner" and "divorced or widowed"), race/color ("white" and "non-white"), land bond ("owner" and "non-owner") and socioeconomic class ("A or B", "C" and "D or E"), according to the Criteria of Economic Classification Brazil. Labor variables were investigated by questioning the current type of production ("conventional" and "non-conventional") and the workload (hours/week) ("less than or equal to 40 h" and "more than 40 h"). Lifestyle variables included alcohol consumption, categorized as "non-drinking" and "drinking"; smoking, assessed according to the Smoker Approach and Treatment Consensus and categorized as "non-smoker" and "current and past smoker"; practice of physical activity extra-field ("yes" or "no"). As a behavioral variable, we used the consumption of the food produced or sold, categorized as "yes" or "no". Among models 1 to 3, the enter method was used. The final model was performed using the Forward LR method. For all of them, the assumptions of absence of multicollinearity and absence of outliers were respected.

Odds ratios (OR), adjusted 95% confidence intervals (CI) and a 5% significance level were presented. All analyzes were conducted in R software (4.0.3) for Windows. The significance level adopted was 5%.

3. Results

The result of the analysis of the ROC curve of the conicity index, according to sex, based on the NCEP criterion can be seen in Figure 1. It is noted that for women, the conicity index has an area under the curve (AUC) corresponding to 0.804 (95% CI: 0.748–0.860, p < 0.001), while in men the AUC was 0.850 (95% CI: 0.787–0.913, p < 0.001) (Figure 1).

Figure 2 shows the result of the ROC curve based on the IDF criterion in both sexes. It is observed that in women, the conicity index has an AUC of 0.784 (95% CI: 0.728–0.840, p < 0.001), while in men it was 0.845 (95% CI: 0.796–0.895, p < 0.001).

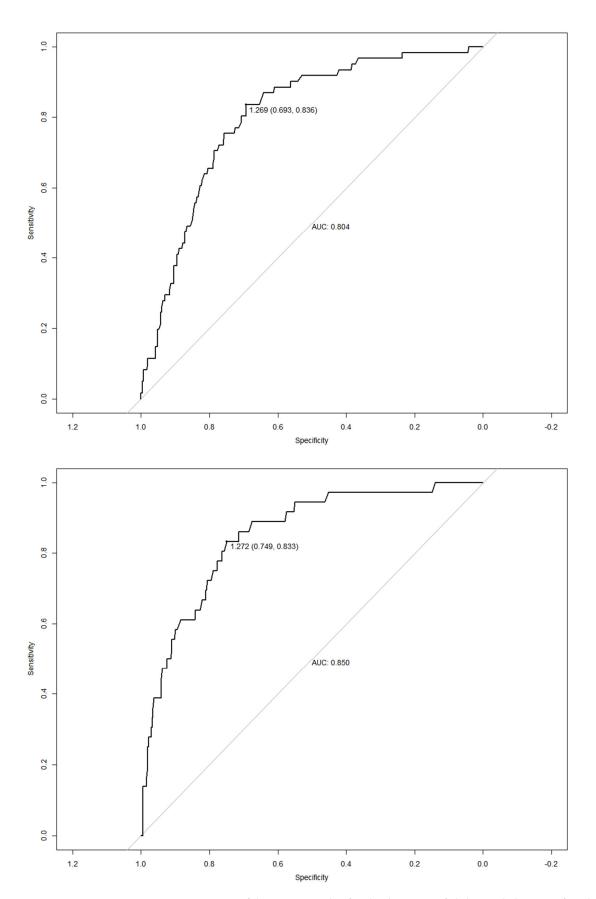


Figure 1. ROC curves of the conicity index for the diagnosis of abdominal obesity in female and male Brazilian rural workers according to the NCEP – ATP III criteria, respectively. AUC—area under the curve.

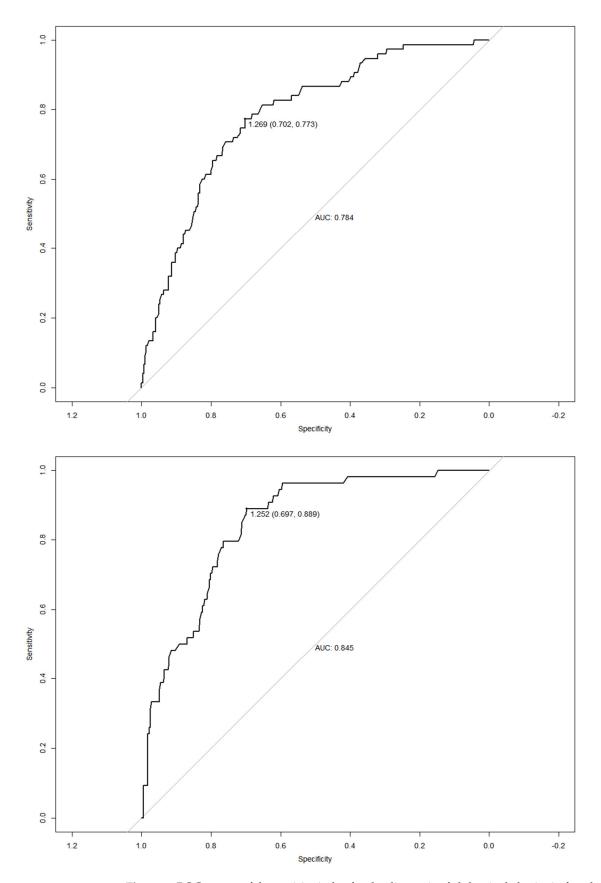


Figure 2. ROC curves of the conicity index for the diagnosis of abdominal obesity in female and male Brazilian rural workers according to the IDF criteria, respectively. AUC—area under the curve.

The cut points were similar in females according to both criteria, resulting in a single cut-off of 1.269. Differences were found in the accuracy criteria: NCEP-ATP III with sensitivity and specificity of 83.60% and 69.30%, respectively, and IDF with sensitivity and specificity of 77.3% and 70.2%, respectively. In males, the cut points showed differences, resulting in 1.272 according to the NCEP-ATP III (sensitivity and specificity of 83.30% and 74.90%, respectively) and 1.252 according to the IDF (sensitivity and specificity of 88, 90% and 69.70%, respectively) (Table 1).

Table 1. Cut points and diagnostic performance measures for the conicity index according to the MetS criteria.

| Variables | NCEP- | -ATP III | IDF | | | |
|------------------|---------------------|---------------------|---------------------|---------------------|--|--|
| | Men (95% CI) | Women (95% CI) | Men (95% CI) | Women (95% CI) | | |
| Cut points | 1.272 | 1.269 | 1.252 | 1.269 | | |
| AUC ¹ | 0.850 (0.787-0.913) | 0.804 (0.748-0.860) | 0.845 (0.796-0.895) | 0.784 (0.728-0.840) | | |
| Accuracy | 0.757 (0.756-0.758) | 0.717 (0.716-0.718) | 0.722 (0.721-0.73) | 0.717 (0.716-0.718 | | |
| Sensitivity | 0.833 (0.712-0.955) | 0.836 (0.743-0.929) | 0.889 (0.805-0.973) | 0.773 (0.679–0.868) | | |
| Specificity | 0.749(0.705-0.793) | 0.693 (0.642-0.744) | 0.697 (0.649-0.745) | 0.702 (0.651-0.754) | | |
| PPV ² | 0.244(0.168-0.320) | 0.347 (0.270-0.424) | 0.310 (0.237-0.382) | 0.395 (0.316-0.474) | | |
| NPV ³ | 0.979(0.962-0.996) | 0.956 (0.929–0.983) | 0.976 (0.957-0.995) | 0.925 (0.891-0.959) | | |

¹ AUC—area under the curve; ² PPV—positive predictive value; ³ NPV—negative predictive value. 95% CI— 95% confidence interval.

Table 2 presents the bivariate analysis of indicators in relation to sociodemographic, lifestyle and occupational variables. For the IDF criterion, there are proportional differences between age (p < 0.001), education (p < 0.001), marital status (p = 0.010), smoking (p = 0.032), physical activity (p = 0.008) and BMI (p < 0.001). In the NCEP criterion, the variables that showed differences were age (p < 0.001), education (p < 0.001), education (p < 0.001), marital status (p = 0.022), alcohol consumption (p = 0.030), physical activity (p = 0.030), physical activity (p = 0.013) and BMI (p < 0.001).

Table 2. Sociodemographic, food consumption and occupational factors of rural works.

| | | IDF | NCEP-ATP III | | | |
|-------------------------------|-----------------------------|----------------|-----------------|----------------|----------------|-----------------|
| Variables | CI ¹ Adequate | CI Elevated | <i>p</i> -Value | CI Adequate | CI Elevated | <i>p</i> -Value |
| Sex | | | 0.836 | | | 0.014 |
| Women | 227 (29.06%) | 147 (18.82%) | | 227 (29.06%) | 147 (18.83%) | |
| Men | 251 (32.13%) | 156 (19.97%) | | 282 (36.10%) | 125 (16.01%) | |
| Age (group) | | | < 0.001 | | | < 0.001 |
| ≤ 29 years | 153 (19.60%) | 23 (2.94%) | | 155 (19.85%) | 21 (2.69%) | |
| \geq 30 to 39 years | 173 (22.15%) | 70 (8.96%) | | 183 (23.45%) | 60 (7.68%) | |
| >40 to 49 years | 98 (12.55%) | 103 (13.19%) | | 108 (13.82%) | 93 (11.90%) | |
| ≥ 50 years or more | 54 (6.91%) | 107 (13.70%) | | 63 (8.06%) | 98 (12.55%) | |
| Schooling (years) | | | < 0.001 | | · · · · | < 0.001 |
| <4 years | 286 (36.62%) | 240 (30.73%) | | 308 (39.44%) | 218 (27.91%) | |
| 4 to 8 years | 129 (16.52%) | 42 (5.37%) | | 132 (16.90%) | 39 (5.00%) | |
| >8 years | 63 (8.06%) | 21 (2.70%) | | 69 (8.83%) | 15 (1.92%) | |
| Marital status | | | 0.010 | | | 0.022 |
| Married or lives with partner | 401 (51.34%) | 270 (34.57%) | | 430 (55.05%) | 241 (30.86%) | |
| Divorced or widowed | 30 (3.84%) | 21 (2.68%) | | 31 (3.96%) | 20 (2.57%) | |
| Unmarried | 47 (6.02%) | 12 (1.53%) | | 48 (6.15%) | 11 (1.41%) | |
| Self-referred race/color | | | 0.281 | | | 0.324 |
| No white | 59 (7.55%) | 29 (3.71%) | | 62 (7.94%) | 26 (3.33%) | |
| White | 419 (53.65%) | 274 (35.08%) | | 447 (57.23%) | 246 (31.50%) | |
| Socioeconomic class | . , | . , | 0.774 | . , | . , | 0.973 |
| Class A or B | 35 (4.48%) | 23 (2.94%) | | 37 (4.74%) | 21 (2.68%) | |
| Class C | 235 (30.09%) | 156 (19.97%) | | 255 (32.65%) | 136 (17.41%) | |

| | | IDF | NCEP-ATP III | | | | |
|-------------------------------------|-----------------------------|----------------|-----------------|--------------------------------------|---|-----------------|--|
| Variables | CI ¹ Adequate | CI Elevated | <i>p</i> -Value | CI Adequate | CI Elevated | <i>p</i> -Value | |
| Class D or E | 208 (26.63%) | 124 (15.87%) | | 217 (27.78%) | 115 (14.72%) | | |
| Land bond | · · · · | · · · · | 0.179 | · · · · | · · · · | 0.324 | |
| Owner | 359 (45.96%) | 241 (30.85%) | | 385 (49.30%) | 215 (27.53%) | | |
| No owner | 119 (15.23%) | 62 (7.94%) | | 124 (15.87%) | 57 (7.30%) | | |
| Type of production | | | 0.112 | | | 0.268 | |
| Conventional | 438 (56.15%) | 267 (34.23%) | | 464 (59.49%) | 241(30.90%) | | |
| No conventional | 39 (5.00%) | 36 (4.61%) | | 44 (5.64%) | 31 (3.97%) | | |
| Workload | | | 0.057 | | | 0.003 | |
| <40 h/week | 87 (11.13%) | 73 (9.34%) | | 88 (11.26%) | 72 (9.22%) | | |
| >40 h/week | 391 (50.06%) | 230 (29.44%) | | 421 (53.90%) | 200 (25.60%) | | |
| Alcohol intake | · · · · · | · · · · | 0.239 | · · · · | () | 0.030 | |
| No | 259 (33.16%) | 178 (22.80%) | | 270 (34.57%) | 167 (21.38%) | | |
| Yes | 219 (28.05%) | 125 (16.00%) | | 239 (30.60%) | 105 (13.44%) | | |
| Smoking | (| (| 0.032 | (, | (, , , , , , , , , , , , , , , , , , , | 0.365 | |
| No smoking | 415 (53.13%) | 245 (31.37%) | | 435 (55.70%) | 225 (28.80%) | | |
| Current smoking or past | 63 (8.06%) | 58 (7.42%) | | 74 (9.47%) | 47 (6.02%) | | |
| Physical activity off field | | | 0.008 | | () | 0.013 | |
| Do not practice | 373 (47.76%) | 263 (33.67%) | | 400 (51.21%) | 236 (30.21%) | | |
| Below recommended | 63 (8.06%) | 23 (2.94%) | | 67 (8.58%) | 19 (2.43%) | | |
| Within the recommended | 42 (5.37%) | 17 (2.17%) | | 42 (5.37%) | 17 (2.17%) | | |
| Body mass index | | | < 0.001 | | | < 0.001 | |
| Low weight/eutrophy | 339 (43.40%) | 44 (5.63%) | | 349 (44.68%) | 34 (4.35%) | | |
| Overweight/obesity | 139 (17.80%) | 259 (33.16%) | | 160 (20.48%) | 238 (30.47%) | | |
| Consumption of minimally processed | (| | 0.476 | (| | 0.795 | |
| Lower contribution | 217 (29.68%) | 148 (20.24%) | | 239 (32.01%) | 127 (17.92%) | | |
| Higher contribution | 228 (31.20%) | 138 (18.87%) | | 234 (32.69%) | 131 (17.37%) | | |
| Consumption of culinary ingredients | () | | 0.118 | | | 0.056 | |
| Lower contribution | 233 (31.87%) | 132 (18.05%) | | 249 (34.06%) | 116 (15.86%) | | |
| Higher contribution | 212 (29.00%) | 154 (21.06%) | | 224 (30.64%) | 142 (19.42%) | | |
| Consumption of processed | (_/ | (| 0.198 | ((((((((((((((((((((((((((((((((((((| (-/-/-// | 0.253 | |
| Lower contribution | 212 (29.00%) | 151 (20.65%) | | 227 (31.05%) | 136 (18.60%) | 0.200 | |
| Higher contribution | 233 (31.87%) | 135 (21.20%) | | 246 (33.65%) | 112 (16.68%) | | |
| Consumption of ultra-processed | | | 0.076 | | (| 0.301 | |
| Lower contribution | 235 (32.14%) | 131 (17.92%) | 0.07.0 | 244 (33.38%) | 122 (16.69%) | 0.001 | |
| Higher contribution | 210 (28.72%) | 155 (21.20%) | | 229 (31.32%) | 136 (18.60%) | | |
| Consumption products you sell | | () | 0.341 | (0110_/0) | 100 (10:00 /0) | 0.098 | |
| No | 22 (2.81%) | 9 (1.15%) | 0.011 | 25 (3.20%) | 6 (0.76%) | 0.070 | |
| Yes | 456 (58.38%) | 294 (37.64%) | | 484 (61.97%) | 266 (34.05%) | | |

Table 2. Cont.

¹ CI—conicity index.

The hierarchical logistic regression between the MetS components according to IDF and the conicity index can be seen in Table 3. It was found that age ≤ 29 years (OR: 11.31; p < 0.001; 95% CI: 5.82–21.98), 30 to 39 years (OR: 4.53; p < 0.001; 95% CI: 2.80–7.34) and 40 to 49 years (OR: 1.82; p = 0.009; 95% CI: 1.15–2.88) and working more than 40 h a week (OR: 1.74; p = 0.009; 95% CI: 1.14–2.67) were associated with the presence of abdominal fat. Regarding the degree of processing of culinary ingredients, the reduction in consumption of these ingredients was 1.66 times more likely to have a metabolic risk (OR: 1.66; p = 0.004; 95% CI: 1.17–2.35).

Table 4 shows the results of the hierarchical logistic regression between the MetS components according to NCEP-ATP III and the conicity index. Age \leq 29 years (OR: 12.51; p < 0.001; 95% CI: 6.46–24.25), 30 to 39 years (OR: 4.75; p < 0.001; 95% CI: 2.94–7.68) and 40 to 49 years (OR: 1.86; p = 0.006; 95% CI: 1.19–2.92) and reduction in consumption of culinary ingredients (OR: 1.57; p = 0.008; 95% CI: 1.12–2.22) were associated with increased abdominal fat, while consuming the products they sell or produce reduced the chances of having abdominal fat by 64% (OR: 0.36; p = 0.040; 95% CI: 0.13–0.95).

| | Model 1 | | Model 2 | | Model 3 | | Final Model | |
|---|-----------------|----------------------------|-----------------|----------------------------|-----------------|----------------------------|-----------------|----------------------------|
| Variables | <i>p</i> -Value | OR (CI _{95%}) |
| Age (group) | | | | | | | | |
| \leq 29 years | < 0.001 | 10.06 (5.51–18.39) | < 0.001 | 9.54 (5.20–17.51) | < 0.001 | 11.31 (5.82–21.98) | < 0.001 | 11.31 (5.82–21.98) |
| \geq 30 to 39 years | < 0.001 | 4.61 (2.95–7.20) | < 0.001 | 4.33 (2.76–6.80) | 0.008 | 4.53 (2.80–7.34) | < 0.001 | 4.53 (2.80–7.34) |
| \geq 40 to 49 years | 0.005 | 1.83 (1.19–2.79) | 0.010 | 1.74 (1.13–2.67) | 0.009 | 1.82 (1.15–2.88) | 0.009 | 1.82 (1.15–2.88) |
| ≥50 years or more Marital status | | 1 | | 1 | | 1 | | 1 |
| Married or lives with partner | | 1 1.27 | | 1 1.34 | | 1 1.21 | | 1 1.21 |
| Unmarried | 0.523 | (0.60–2.64) 1.43 | 0.431 | (0.64–2.82) 1.52 | 0.618 | (0.56–2.57) 1.51 | 0.618 | (0.56–2.57) 1.51 |
| Divorced or widowed | 0.256 | (0.77–2.66) | 0.186 | (0.81–2.86) | 0.201 | (0.80–2.85) | 0.201 | (0.80–2.85) |
| Schooling (years) <4 years | | 1 | | 1 | | 1 | | 1 |
| 4 to 8 years | 0.540 | 1.15 (0.73–1.80) | 0.465 | 1.18 (0.75–1.86) | 0.608 | 1.13 (0.70–1.82) | 0.608 | 1.13 (0.70–1.82) |
| >8 years | 0.066 | 1.81 (0.96–3.41) | 0.054 | 1.87 (0.98–3.56) | 0.053 | 1.92 (0.99–3.73) | 0.053 | 1.92 (0.99–3.73) |
| Bond with the earth No owner | | 1 | | 1 | | 1 | | 1 |
| Owner | 0.636 | 0.91 (0.61–1.34) | 0.413 | 0.84 (0.56–1.26) | 0.448 | 0.84 (0.55–1.29) | 0.448 | 0.84 (0.55–1.29) |
| Workload $\leq 40 \text{ h/week}$ | | | | 1 | | 1 | | 1 |
| >40 h/week | | | 0.007 | 1.72 (1.15–2.56) | 0.009 | 1.74 (1.14–2.67) | 0.009 | 1.74 (1.14–2.67) |
| Type of production Conventional | | | | 1 | | 1 | | 1 |
| No conventional | | | 0.537 | 0.84 (0.49–1.44) | 0.374 | 0.77 (0.44–1.35) | 0.374 | 0.77 (0.44–1.35) |
| Physical activity off field Do not practice | | | | (0.1)-1.11) | | 1 | | 1 |
| Below recommended | | | | | 0.282 | 1.38 (0.76–2.52) | 0.282 | 1.38 (0.76–2.52) |
| Within the recommended | | | | | 0.830 | 1.07 (0.54–2.14) | 0.830 | 1.07 (0.54–2.14) |
| Smoking No smoking | | | | | | 1 1.23 | | 1 1.23 |
| Current smoking or past | | | | | 0.382 | (0.76–1.99) | 0.382 | (0.76–1.99) |
| Degree of processing Culinary ingredients Higher contribution | | | | | | 1 | | 1 |
| Lower contribution | | | | | 0.004 | 1.66 (1.17–2.35) | 0.004 | 1.66 (1.17–2.35) |
| Processed Higher contribution | | | | | | 1 | | 1 |
| Lower contribution | | | | | 0.759 | 0.94 (0.65–1.35) | 0.759 | 0.94 (0.65–1.35) |
| Ultra-processed Higher contribution | | | | | | (0.65-1.55) | | (0.65–1.55) |
| Lower contribution | | | | | 0.245 | 1.24 (0.86–1.79) | 0.245 | 1.24 (0.86–1.79) |

Table 3. Hierarchical logistic regression between the associated variables in the bivariate analysis and the conicity index according to the components of the IDF.

| | Model 1 | | Model 2 | | Model 3 | | Final Model | |
|---|-----------------|----------------------------|-----------------|----------------------------|-----------------|----------------------------|-----------------|----------------------------|
| Variables | <i>p</i> -Value | OR (CI _{95%}) |
| Sex | | | | | | | | |
| Women | | 1 | | 1 | | 1 | | 1 |
| Men | 0.001 | 1.68 (1.21–2.33) | 0.010 | 1.55 (1.10–2.18) | 0.074 | 1.41 (0.96–2.07) | 0.074 | 1.41 (0.96–2.07) |
| Age (group) | | 11.20 | | 10.00 | | 10 51 | | 10 51 |
| \leq 29 years | < 0.001 | 11.39 (6.18–20.98) | < 0.001 | 10.98 (5.95–20.27) | < 0.001 | 12.51 (6.46–24.25) | < 0.001 | 12.51 (6.46–24.25) |
| \geq 30 to 39 years | < 0.001 | 4.93 (3.14–7.76) | < 0.001 | 4.73 (3.00–7.46) | < 0.001 | 4.75 (2.94–7.68) | < 0.001 | 4.75 (2.94–7.68) |
| \geq 40 to 49 years | 0.003 | 1.90 (1.24–2.92) | 0.004 | 1.85 (1.21–2.85) | 0.006 | 1.86 (1.19–2.92) | 0.006 | 1.86 (1.19–2.92) |
| ≥50 years or more Marital status | | 1 | | 1 | | 1 | | 1 |
| Married or lives with partner | | 1 | | 1 | | | 1 | |
| Unmarried | 0.795 | 1.10 (0.52–2.34) | 0.681 | 1.17 (0.55–2.49) | 0.865 | 1.06 (0.49–2.31) | 0.865 | 1.06 (0.49–2.31) |
| Divorced or widowed | 0.147 | 1.58 (0.84–2.96) | 0.123 | 1.64 (0.87–3.07) | 0.157 | 1.58 (0.83–2.99) | 0.157 | 1.58 (0.83–2.99) |
| Schooling (years) <4 years | | 1 | | 1 | | 1 | | 1 |
| 4 to 8 years | 0.645 | 1.11 (0.70–1.74) | 0.607 | 1.12 (0.71–1.77) | 0.679 | 1.10 (0.68–1.78) | 0.679 | 1.10 (0.68–1.78) |
| >8 years | 0.083 | 1.74 (0.92–3.29) | 0.079 | 1.76 (0.93–3.32) | 0.126 | 1.66 (0.86–3.19) | 0.126 | 1.66 (0.86–3.19) |
| Workload ≤40 h/week | | | | 1 | | 1 | | 1 |
| >40 h/week | | | 0.088 | 1.42 (0.94–2.13) | 0.055 | 1.53 (0.98–2.36) | 0.055 | 1.53 (1.01–2.36) |
| Physical activity off field Do not practice | | | | (001 200) | | 1 | | 1 |
| Below recommended | | | | | 0.302 | 1.37 (0.75–2.50) | 0.302 | 1.37 (0.75–2.50) |
| Within the recommended | | | | | 0.848 | 1.06 (0.53–2.13) | 0.848 | 1.06 (0.53–2.13) |
| Alcohol intake No | | | | | | 1 | | 1 |
| Yes | | | | | 0.679 | 1.08 (0.74–1.56) | 0.679 | 1.08 (0.74–1.56) |
| Degree of processing Culinary ingredients Higher contribution | | | | | | 1 | | 1 |
| Lower contribution | | | | | 0.008 | 1.57 (1.12–2.22) | 0.008 | 1.57 (1.12–2.22) |
| Consumption products you sell No | | | | | | (1.12-2.22) | | (1.12-2.22) |
| Yes | | | | | 0.040 | 0.36 (0.13–0.95) | 0.040 | 0.36 (0.13–0.95) |

Table 4. Hierarchical logistic regression between the associated variables in the bivariate analysis and the conicity index according to the components of the NCEP-ATP III.

OR—odds ratio; 95% CI—confidence interval 95%.

4. Discussion

This study is a pioneer in identifying the cut-off point and showing the effectiveness and reliability of the obesity index as evidence of MetS stratified by sex in rural workers. The cut-off points found can be used in epidemiological studies and in clinical practice in this specific population.

The conicity index was proposed by Valdez [15] as an anthropometric indicator to identify the distribution of body fat, specifically abdominal adiposity. However, it presents difficulties in its routine use due to the absence of standardized cut-off points for reference in the population. According to the mathematical equation that gives rise to the index, the denominator determines the shape of the body by considering the height and weight

of individuals. This is an advantage over other anthropometric indicators, as it allows comparison between different groups [15].

The cut point of the conicity index in Brazilians most used in the literature was performed in healthy adults to discriminate the high coronary risk, presenting values of 1.25 and 1.18 for men and women, respectively [30]. In addition to this study, other authors identified high values when they used the diagnostic criteria for MetS, between 1.61 and 1.56 for men and women, respectively [31], and 1.36 for women [32]. When compared with the values found in this study, a divergence is observed mainly in females, highlighting the need to use the specific cut point for rural workers.

Regarding age, both diagnostic criteria (IDF and NCEP-ATP III) showed results that indicate a higher risk for the presence of abdominal fat in younger individuals. This finding is already consolidated in the literature, showing that younger individuals have greater general and abdominal obesity compared to the elderly [33], and in rural areas it was shown that the younger population tended to consume more ultra-processed foods [26].

The NOVA classification system was first presented by Monteiro et al. [34] and classifies all foods into four groups according to the nature, extent and purpose of industrial food processing, namely: Group 1—foods unprocessed/minimally processed; Group 2—processed culinary ingredients; Group 3—processed foods; Group 4—ultra-processed foods [34]. Regarding culinary ingredients, these are used to season and cook food and create culinary preparations [22]. The reduction in the consumption of foods with a low degree of processing is a parameter that confers a risk of abdominal obesity above normal.

With regard to our findings, as per the IDF and NCEP classification, reducing the consumption of culinary ingredients increases the risk of high abdominal fat by 1.66-fold and 1.57-fold, respectively. This finding is in line with the literature since other studies, including a study carried out with this same population, found that lower adherence to traditional dietary patterns is associated with a higher prevalence of both general obesity and abdominal obesity [11,35–37]. Historically, the food pattern of rural populations consists of the consumption of rice, beans, cassava flour, coffee, cow's milk and bread, but in recent years there has been an increase in the consumption of traditional industrialized products in urban areas, which highlights a new scenario with the influence of globalization [26,38–40].

In the young adult population, Santana et al. [41] relate the consumption of ultraprocessed foods to the presence of abdominal obesity due to their poor nutritional quality and higher total energy value, which is one of the cardiometabolic risk factors in the progressive increase in these individuals. Concomitantly with obesity, there has also been an increase in the development of hypertension, diabetes and cardiovascular diseases, chronic non-communicable diseases typical of a diet rich in industrialized foods that are the result of metabolic abnormalities, often caused by inflammation resulting from excess sugars and fats in ultra-processed foods [4,35,42,43].

In addition, this study showed that the consumption of food for sale is considered a factor of abdominal obesity. It is already established in obesity that fresh foods are less energetically dense and are rich in vitamins, minerals and fibers that act as protectors against the literature and cardiometabolic diseases [44–46]. It is noteworthy that among the main agricultural productions in this region are the planting of potatoes, beans, cassava, corn, tomatoes, bananas, coffee, guava, lemons, grapes, passion fruit and palm hearts [47] and many of these represent healthy choices important for this population. In a study carried out with the same sample, it was found that more than half of the population (64.9%) consumed tomatoes and that bananas and lemons were the fruits most consumed by the participants [6].

The limitation of the present study is its cross-sectional nature, which requires greater caution in interpreting the results due to the possibility of reverse causality. Despite this, the importance of the findings presented is justified, since the study showed high rigor in the sampling process, reaching a representative sample of the population of farmers. The lack of articles that use the same methodology in rural populations represents both a limitation and a strong point of the work.

It represents a limitation as it makes it difficult to compare the results, but it represents a strong point, because, in addition to its unprecedented nature, it may represent the beginning of new research aimed at the health of this population, especially the assessment of body adiposity, in order to reach a consensus on the future cut-off point and publicize the use of this technique in clinical practice.

5. Conclusions

In conclusion, cut-off points of 1.269 for women and 1.252 to 1.272 for men for the conicity index showed good sensitivity and specificity for identifying abdominal obesity in farmers. Thus, the conicity index proves to be an accurate and easily accessible marker to be incorporated into clinical practice in this population.

There is a need to improve eating habits and promote healthier eating environments for individuals, while respecting the traditional food culture, especially to contain the advancement of MetS in rural areas. Finally, due to the multicausal nature of MetS, coping strategies for this condition must include a multiple, intersectoral and interdisciplinary approach.

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References

- Cornier, M.A.; Dabelea, D.; Hernandez, T.L.; Lindstrom, R.C.; Steig, A.J.; Stob, N.R.; Van Pelt, R.E.; Wang, H.; Eckel, R.H. The metabolic syndrome. *Endocr. Rev.* 2008, 7, 777–822. [CrossRef] [PubMed]
- Oliveira, L.V.; Santos, B.N.; Machdo, I.E.; Malta, D.C.; Velasquez-Melendez, G.; Felisbino-Mendes, M.S. Prevalência da Síndrome Metabólica e seus componentes na população adulta brasileira. *Ciênc. Saúde Coletiva* 2020, 25, 4269–4280. [CrossRef]
- 3. Lourenço, A.E. The meaning of 'rural' in rural health: A review and case study from Brazil. *Glob. Public Health* **2012**, *7*, 1–13. [CrossRef] [PubMed]
- 4. Luz, T.C.D.; Cattafesta, M.; Petarli, G.B.; Meneghetti, J.P.; Zandonade, E.; Bezerra, O.M.D.P.A.; Salaroli, L.B. Cardiovascular risk factors in a Brazilian rural population. *Ciênc. Saúde Coletiva* **2020**, *25*, 3921–3932. [CrossRef]
- 5. Ferreira, J.R.S.; Zandonade, E.; Bezerra, O.M.P.; Salaroli, L.B. Insulin resistance by the triglyceride-glucose index in a rural Brazilian population. *Arch. Endocrinol. Metab.* **2022**, 1–8. [CrossRef]
- 6. Cattafesta, M.; Petarli, G.B.; da Luz, T.C.; Zandonade, E.; Bezerra, O.M.P.; Salaroli, L.B. Dietary patterns of Brazilian farmers and their relation with sociodemographic, labor, and lifestyle conditions. *Nutr. J.* **2020**, *19*, 23. [CrossRef]
- Powell-Wiley, T.M.; Poirier, P.; Burke, L.E.; Després, J.P.; Gordon-Larsen, P.; Lavie, C.J.; Lear, S.A.; Ndumele, C.E.; Neeland, I.J.; Sanders, P.; et al. Obesity and Cardiovascular Disease: A Scientific Statement from the American Heart Association. *Circulation* 2021, 143, e984–e1010. [CrossRef]
- 8. Rychter, A.M.; Ratajczak, A.E.; Zawada, A.; Dobrowolska, A.; Krela-Kaźmierczak, I. Non-systematic review of diet and nutritional risk factors of cardiovascular disease in obesity. *Nutrients* **2020**, *12*, 814. [CrossRef]

- Quaye, L.; Owiredu, W.K.B.A.; Amidu, N.; Dapare, P.P.M.; Adams, Y. Comparative Abilities of Body Mass Index, Waist Circumference, Abdominal Volume Index, Body Adiposity Index, and Conicity Index as Predictive Screening Tools for Metabolic Syndrome among Apparently Healthy Ghanaian Adults. J. Obes. 2019, 2019, 8143179. [CrossRef] [PubMed]
- 10. Koliaki, C.; Liatis, S.; Kokkinos, A. Obesity and cardiovascular disease: Revisiting an old relationship. *Metabolism* **2019**, *92*, 98–107. [CrossRef] [PubMed]
- 11. Cattafesta, M.; Petarli, G.B.; Zandonade, E.; Bezerra, O.M.P.A.; Abreu, S.M.R.; Salaroli, L.B. Prevalence and determinants of obesity and abdominal obesity among rural workers in Southeastern Brazil. *PLoS ONE* **2022**, *17*, e0270233. [CrossRef] [PubMed]
- Su, T.T.; Amiri, M.; Mohd Hairi, F.; Thangiah, N.; Dahlui, M.; Majid, H.A. Body composition indices and predicted cardiovascular disease risk profile among urban dwellers in Malaysia. *Biomed. Res. Int.* 2015, 2015, 174821. [CrossRef] [PubMed]
- 13. Rahman, M.; Berenson, A.B. Accuracy of current body mass index obesity classification for white, black and Hispanic reproductiveage women. *Obstet. Gynecol.* 2010, 115, 982. [CrossRef] [PubMed]
- 14. Nevill, A.M.; Stewart, A.D.; Olds, T.; Holder, R. Relationship between adiposity and body size reveals limitations of BMI. *Am. J. Phys. Anthropol* **2006**, *129*, 151–156. [CrossRef]
- 15. Valdez, R. A simple model-based index of abdominal adiposity. J. Clin. Epidemiol. 1991, 44, 955–956. [CrossRef]
- Motamed, N.; Sohrabi, M.; Poustchi, H.; Maadi, M.; Malek, M.; Keyvani, H.; Amoli, M.S.; Zamani, F. The six obesity indices, which one is more compatible with metabolic syndrome? A population-based study. *Diabetes Metab. Syndr.* 2016, 11, 173–177. [CrossRef]
- 17. Wang, H.; Liu, A.; Zhao, T.; Gong, X.; Pang, T.; Zhou, Y.; Xiao, Y.; Yan, Y.; Fan, C.; Teng, W.; et al. Comparison of anthropometric indices for predicting the risk of metabolic syndrome and its components in Chinese adults: A prospective, longitudinal study. *BMJ Open* **2017**, *7*, e016062. [CrossRef]
- 18. Petarli, G.B.; Cattafesta, M.; Sant'Anna, M.M.; Bezerra, O.M.d.P.A.; Zandonade, E.; Salaroli, L.B. Multimorbidity and complex multimorbidity in Brazilian rural workers. *PLoS ONE* **2019**, *14*, e0225416. [CrossRef]
- 19. Sociedade Brasileira de Cardiologia; Sociedade Brasileira de Nefrologia. VI Brazilian guidelines on hypertension. *Arq. Bras. De Cardiol.* **2010**, *95*, 1–51.
- 20. Willett, W. Nutritional Epidemiology, 1st ed.; Oxford University Press: Oxford, UK, 2012.
- 21. Nusser, S.M.; Carriquiry, A.L.; Fuller, W.A. A semi parametric transformation approach to estimating usual intake distributions. *CARD Work. Pap.* **1996**, 1440–1449.
- 22. Brasil Ministério da Saúde. Guia Alimentar Para a População Brasileira, 2nd ed.; Brasil Ministério da Saúde: Brasilia, Brazil, 2014.
- Monteiro, C.A.; Levy, R.B.; Claro, R.M.; Castro, I.R.R.D.; Cannon, G. A new classification of foods based on the extent and purpose of their processing. *Cad. De Saude Publica* 2010, 26, 2039–2049. [CrossRef]
- 24. Monteiro, C.A.; Cannon, G.; Levy, R.; Moubarac, J.C.; Jaime, P.; Martins, A.P.; Canella, D.; Louzada, M.L.; Parra, D.; Ricardo, C.; et al. Classificação dos alimentos. Saúde Pública. NOVA. A estrela brilha. *Word Nutr.* **2016**, *7*, 28–40.
- Monteiro, C.; Cannon, G.; Levy, R.B.; Claro, R.; Moubarac, J.C.; Martins, A.P.; Louzada, M.L.; Baraldi, L.; Canella, D. The food system. Ultra-processing: The big issue for nutrition, disease, health, well-being. *World Nutr.* 2012, *3*, 527–569.
- 26. Cattafesta, M.; Petarli, G.B.; Zandonade, E.; Bezerra, O.M.P.A.; Abreu, S.M.R.; Salaroli, L.B. Energy contribution of NOVA food groups and the nutritional profile of the Brazilian rural workers' diets. *PLoS ONE* **2020**, *28*, e0240756. [CrossRef]
- 27. Lohman, T.G.; Roche, A.F.; Martorell, R. Anthropometric Standardization Reference Manual; Human Kinetics: Champaign, IL, USA, 1988.
- World Health Organization. Obesity: Preventing and Managing the Global Epidemic: Report of a WHO Consultation; WHO Consultation on Obesity: Geneva, Switzerland, 2000; 252p. Available online: http://www.who.int/entity/nutrition/publications/obesity/ WHO_TRS_894/en/index.html (accessed on 2 October 2022).
- 29. de Carvalho, M.H. I Diretriz brasileira de diagnóstico e tratamento da síndrome metabólica. Arq. Bras. De Cardiol. 2005, 84, 1–28.
- 30. Pitanga, F.J.G.; Lessa, I. Sensibilidade e especificidade do índice de conicidade como discriminador do risco coronariano de adultos em Salvador, Brasil. *Rev. Bras. Epidemiol.* **2004**, *7*, 259–269. [CrossRef]
- Ceolin, J.; Engroff, P.; Mattiello, R.; Schwanke, C.H.A. Performance of Anthropometric Indicators in the Prediction of Metabolic Syndrome in the Elderly. *Metab. Syndr. Relat. Disord.* 2019, 17, 232–239. [CrossRef] [PubMed]
- 32. Paula, H.A.A.; Ribeiro, R.C.L.; Rosado, L.E.F.P.L.; Pereira, R.S.F.; Franscechini, F.C.C. Comparação de diferentes critérios de definição para diagnóstico de síndrome metabólica em idosas. *Arq. Bras. Cardiol.* **2010**, *95*, 346–353. [CrossRef] [PubMed]
- Lovsletten, O.; Jacobsen, B.K.; Grimsgaard, S.; Njølstad, I.; Wilsgaard, T.; Løchen, M.L.; Eggen, A.E.; Hopstock, L.A. Prevalence of general and abdominal obesity in 2015–2016 and 8-year longitudinal weight and waist circumference changes in adults and elderly: The Tromsø Study. *BMJ Open* 2020, 10, e038465. [CrossRef]
- 34. Monteiro, C.A.; Cannon, G.; Moubarac, J.C.; Levy, R.B.; Louzada, M.L.C.; Jaime, P.C. The UN Decade of Nutrition, the NOVA food classification and the trouble with ultra-processing. *Public Health Nutr.* **2018**, *21*, 5–17. [CrossRef]
- Martins-Silva, T.; Vaz, J.S.; Mola, C.L.; Assunção, M.C.F.; Tovo-Rodrigues, L. Prevalence of obesity in rural and urban areas in Brazil: National Health Survey, 2013. *Rev. Bras. Epidemiol.* 2019, 22, e190049. [CrossRef] [PubMed]
- Abeywickrama, H.M.; Wimalasiri, K.M.S.; Koyama, Y.; Uchiyama, M.; Shimizu, U.; Chandrajith, R.; Nanayakkara, N. Assessment of Nutritional Status and Dietary Pattern of a Rural Adult Population in Dry Zone, Sri Lanka. *Int. J. Environ. Res. Public Health* 2020, 17, 150. [CrossRef] [PubMed]
- NCD Risk Factor Collaboration (NCD-RisC). Rising rural body-mass index is the main driver of the global obesity epidemic in adults. *Nature* 2019, 569, 260–264. [CrossRef] [PubMed]

- Carvalho, E.O.; Da Rocha, E.F. Consumo para alimentação da população adulta residente na zona rural do município de Ibatiba (ES, Brasil). *Ciênc. Saúde Coletiva* 2011, 16, 179. [CrossRef]
- Canella, D.S.; Louzada, M.L.C.; Claro, R.M.; Costa, J.C.; Bandoni, D.H.; Levy, R.B.; Martins, A.P.B. Consumption of vegetables and their relation with ultra-processed foods in Brazil. *Rev. Saúde Pública* 2018, 52, 50. [CrossRef]
- 40. Fernandes, M.P.; Bielemann, R.M.; Fassa, A.C.G. Factors associated with the quality of the diet of residents of a rural area in Southern Brazil. *Rev. Saúde Pública* 2018, 52, 6s. [CrossRef]
- Santana, G.J.; Silva, N.J.; Costa, J.O.; Vásquez, C.M.P.; Vila-Nova, T.M.S.; Vieira, D.A.S.; Pires, L.V.; Fagundes, A.A.; Barbosa, K.B.F. Contribution of minimally processed and ultra-processed foods to the cardiometabolic risk of Brazilian young adults: A cross-sectional study. *Nutr. Hosp.* 2021, 38, 328–336. [CrossRef]
- 42. Salaroli, L.B.; Cattafesta, M.; Petarli, G.B.; Ribeiro, S.A.V.; Soares, A.C.O.; Zandonade, E.; Bezerra, O.M.P.A.; Mill, J.G. Prevalence and factors associated with arterial hypertension in a Brazilian rural working population. *Clinics* **2020**, *75*, e1603. [CrossRef]
- Bolte, L.A.; Vich Vila, A.; Imhann, F.; Collij, V.; Gacesa, R.; Peters, V.; Wijmenga, C.; Kurilshikov, A.; Campmans-Kuijpers, M.J.E.; Fu, J.; et al. Long-term dietary patterns are associated with pro-inflammatory and anti-inflammatory features of the gut microbiome. *Gut* 2021, 70, 1287–1298. [CrossRef]
- 44. Cena, H.; Calder, P.C. Defining a healthy diet: Evidence for the role of contemporary dietary patterns in health and disease. *Nutrients* **2020**, *12*, 334. [CrossRef]
- Chiavaroli, L.; Viguiliouk, E.; Nishi, S.K.; Blanco Mejia, S.; Rahelić, D.; Kahleová, H.; Salas-Salvadó, J.; Kendall, C.W.; Sievenpiper, J.L. DASH dietary pattern and cardiometabolic outcomes: An umbrella review of systematic reviews and meta-analyses. *Nutrients* 2019, 11, 338. [CrossRef] [PubMed]
- 46. Fraga, C.G.; Croft, K.D.; Kennedy, D.O.; Tomás-Barberán, F.A. The effects of polyphenols and other bioactives on human health. *Food Funct.* **2019**, *10*, 514–528. [CrossRef] [PubMed]
- Intituto Brasileiro de Geografia e Estatística. Produção Agrícola, lavoura Permanente. Santa Maria de Jetibá, Espírito Santo. 2021. Available online: https://cidades.ibge.gov.br/brasil/es/santa-maria-de-jetiba/pesquisa/15/11863 (accessed on 10 December 2021).